

Notes on Fundamental Inequalities

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October 2005

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1 Moments

Given a measure space $(\Omega, \mathcal{F}, \nu)$ define the Expected Value as

$$E[h(x)] = \int_{\Omega} h(X(\omega)) d\nu(\omega) \quad (1)$$

We define the following:

- k-th moment of X :

$$E[X^k] \quad \text{with } k \geq 1, k \in \mathbb{N} \quad (2)$$

- k-th absolute moment:

$$E[|X|^k] \quad \text{with } k \geq 0, k \in \mathbb{N} \quad (3)$$

- norm of order k:

$$\|X\|_k = (E[|X|^k])^{1/k} \quad \text{with } k \geq 0, k \in \mathbb{N} \quad (4)$$

- k-th central moment:

$$\mu_k = E[(X - \mu)^k] \quad \text{with } k \geq 1, k \in \mathbb{N} \quad (5)$$

where $\mu = E[X^1]$ (the mean) is the moment of order 1.

Moments and Norms are defined as long as the correspondent integrals exist.

PROPOSITION:

Given $\nu(\Omega) < \infty$, if $E[|X|^r]$ exists and is finite then also $E[|X|^s]$ is finite $\forall 0 \leq s \leq r$. Further on, also $E[X^k]$ exists and it's finite $\forall 0 \leq k \leq r$.

Proof:

To prove the proposition we need the following disequality:

$$|x|^s \leq 1 + |x|^r \quad \forall 0 \leq s \leq r \quad r, s \in \mathbb{N} \quad (6)$$

w.l.g. we consider only positive values of x (the absolute value is a symmetric function).

$$\text{a) } 0 \leq x \leq 1$$

$$\implies 0 \leq x^s \leq 1 \quad \forall s$$

$$\implies x^s \leq 1 + A \quad \text{where } A \text{ is any positive quantity}$$

$$\implies x^s \leq 1 + |x|^r, \quad 0 \leq r \leq s$$

$$\text{b) } x > 1$$

$$\implies x^t \geq 1 \text{ where } t = \frac{r}{s} \geq 1$$

$$\implies x^r \geq x^s$$

$$\implies 1 + x^r \geq x^s$$

Therefore inequality (6) is proved.

Apply the Expected Value to both sides of inequality (6):

$$E[|X|^s] \leq 1 + E[|X|^r] \quad \forall 0 < s \leq r \quad (7)$$

The inequality is maintained since the Expected Value is a linear function. From (7), if $E[|X|^r]$ exists and is finite, also $E[|X|^s]$ is finite. Q.E.D.

The proof of the second part of the proposition is omitted.

EXAMPLE:

In a probability space ($\nu(\Omega) = 1$)

$$\sigma^2 = E[(X - \mu)^2] < \infty \iff E[X^2] < \infty$$

Proof:

\implies

It's easy to prove that $\sigma^2 = E[X^2] - \mu^2$.

$$0 < \sigma^2 < \infty \iff 0 < E[X^2] - \mu^2 < \infty$$

Since μ^2 is positive, $E[X^2]$ must be finite.

←

By the proposition, $E[X^2] < \infty \implies E[X^1] = \mu < \infty$. Since $\sigma^2 = E[X^2] - E[X]^2$ the result follows.

In the following sections we will limit our attention to the case $\nu(\Omega) = 1$, i.e. we restrict ourselves to probability spaces.

2 Base (Generalized Chebyshev) Inequality

Let $g \geq 0$ be an even, non-decreasing function in \mathbb{R}^+ . Let $(\Omega, \mathcal{F}, \nu)$, $\nu(\Omega) = 1$ be a probability space. If g is ν -measurable

$$P(|X| \geq \lambda) \leq \frac{E[g(X)]}{g(\lambda)} \quad \forall \lambda > 0 \quad (8)$$

Proof:

$$\begin{aligned} E[g(X)] &= \int_{\Omega} g(x) d\nu(x) \\ &= \int_{|x| \geq \lambda} g(x) d\nu(x) + \int_{|x| < \lambda} g(x) d\nu(x) \\ &\geq \int_{|x| \geq \lambda} g(x) d\nu(x) \\ &\geq g(\lambda) \int_{|x| \geq \lambda} 1 d\nu(x) \\ &= g(\lambda) P(|X| \geq \lambda) \end{aligned}$$

Q.E.D.

The only non-trivial step is the third, that is given by the fact that g is positive and non-decreasing.

3 Markov Inequality

$$P(|X| \geq \lambda) \leq \frac{E[|X|^r]}{\lambda^r} \quad \forall \lambda > 0 \quad (9)$$

Proof:

The result follows from the Base inequality setting $g : x \mapsto |x|^r$.

4 Chebyshev Inequality

$$P(|X - \mu| \geq \lambda) \leq \frac{\sigma^2}{\lambda^2} \quad \forall \lambda > 0 \quad (10)$$

Proof:

Remember that $\mu = E[X]$ e $\sigma^2 = E[(X - \mu)^2]$.

The result follows from the Base inequality setting $g : (x - \mu) \mapsto (x - \mu)^2$

5 Convex Functions

- A real valued function $f : I \mapsto \mathbb{R}$ is convex if

$$\begin{aligned} f(\alpha x + (1 - \alpha)y) &\leq \alpha f(x) + (1 - \alpha)f(y) \\ \forall x, y \in I \\ \forall \alpha \in [0, 1] \end{aligned} \quad (11)$$

- If f is a continuous function, an equivalent definition is

$$\forall c \in I, \exists b(\cdot) : f(z) \geq f(c) + b(z - c) \quad \forall z \in I$$

Note that the function b is not restricted, can be either a positive or a negative-valued function.

- If f is also differentiable, an equivalent definition is $f'(x)$ is an increasing function in x .
- If f is twice differentiable, f is convex is equivalent to $f''(x) \geq 0 \forall x \in I$

THEOREM: Supporting Hyperplane

Let $f : I \mapsto \mathbb{R}$ be a convex function. Then $\forall x_0 \in I$ exists a linear function $l(x) = a_0 + b_0x$ such that $f(x) \geq l(x) \forall x \neq x_0$, and $f(x_0) = l(x_0)$. The function $l(x)$ is called supporting hyperplane¹.

¹it's called hyperplane because the theorem holds also for convex functions in \mathbb{R}^n

6 Jensen Inequality

Given a probability space $(\Omega, \mathcal{F}, \nu)$, $\nu(\Omega) = 1$, let $g : (a, b) \mapsto \mathbb{R}$ be a convex ν -measurable function, where $-\infty \leq a < b \leq +\infty$.

If $E[X]$ exists and it's finite (and therefore $E[X] \in (a, b)$), then

$$g(E[X]) \leq E[g(X)] \quad (12)$$

For strictly convex functions (12) holds with equality iff $X = E[X]$ *a.e.*

Proof:

Let $l(\cdot)$ be the supporting hyperplane of $g(\cdot)$ in $x_0 \equiv E[X]$. Then

$$\begin{aligned} E[g(X)] &\geq E[l(X)] \\ &= l(E[X]) \quad \text{since } l \text{ is a linear function and } \nu(\Omega) = 1 \\ &= g(E[X]) \quad \text{since } g(\cdot) = l(\cdot) \text{ in } x_0 \equiv E[X] \end{aligned}$$

$g(X) - l(X) \geq 0$, and therefore $E[g(X)] = E[l(X)]$ iff $g(X) = l(X)$ *a.e.* $\iff X = E[X]$ *a.e.*

Q.E.D.

7 Cauchy–Schwarz Inequality

If the $E[X^2]$, $E[Y^2]$ exist and are finite, then also $E[|XY|]$ exists and is finite. Moreover,

$$E[|XY|]^2 \leq E[X^2] E[Y^2] \quad (13)$$

Proof:

(a) Existence:

$$\begin{aligned} (X - Y)^2 &\geq 0 \\ X^2 - 2|XY| + Y^2 &\geq 0 \\ 2|XY| &\leq X^2 + Y^2 \\ |XY| &\leq X^2 + Y^2 \\ E[|XY|] &\leq E[X^2] + E[Y^2] \end{aligned}$$

The transformed r.v. $|XY|$ can only assume positive values, therefore $E[|XY|] = \int_{\Omega} |XY(\omega)| d\nu(\omega)$ exists. From the last inequality (Triangular Inequality), finiteness of $E[X^2]$ and $E[Y^2]$ imply finiteness of $E[|XY|]$. By the proposition proved in the first section, existence and finiteness of $E[|XY|]$ imply existence and finiteness for $E[XY]$

(b) Chauchy–Schwartz inequality:

Given that $E[|XY|]^2 = E[XY]^2$, we prove $E[XY]^2 \leq E[X^2] E[Y^2]$.

$\forall z \in \mathbb{R}$ we have

$$0 \leq E[(zX + Y)^2] = z^2 E[X^2] + 2zE[XY] + E[Y^2]$$

The last equality comes from the linearity of $E[\cdot]$. The discriminant of the quadratic equation in z is

$$\sqrt{4(E[XY]^2 - E[X^2]E[Y^2])},$$

And since the equation is always non–negative,

$$E[XY]^2 - E[X^2]E[Y^2] \leq 0$$

Q.E.D.

EXAMPLE: Inequalities for correlation coefficient

If X, Y are r.v. with finite (and positive) variance, then

$$-1 \leq \rho \leq 1 \quad \text{where } \rho = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)\text{Var}(Y)}}$$

Proof:

Let $X' = X - E[X]$, $Y' = Y - E[Y]$, and apply the Chauchy–Schwartz inequality to the transformed variables:

$$\begin{aligned} E\left[|(X - E[X])(Y - E[Y])|\right]^2 &\leq E[(X - E[X])^2] E[(Y - E[Y])^2] \\ E\left[|(X - E[X])(Y - E[Y])|\right] &\leq \sqrt{\text{Var}(X)\text{Var}(Y)} \end{aligned}$$

$|\cdot|$ is a convex function, by Jensen inequality $|E[Z]| \leq E[|Z|]$:

$$\left| E[(X-E[X])(Y-E[Y])] \right| \leq E \left[|(X-E[X])(Y-E[Y])| \right] \leq \sqrt{\text{Var}(X)\text{Var}(Y)}$$

$$|\text{Cov}(X, Y)| \leq \sqrt{\text{Var}(X)\text{Var}(Y)}$$

$$|\rho| \leq 1$$

$$-1 \leq \rho \leq 1$$

Q.E.D.